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Remittances and the Dutch Disease in Sub-Saharan Africa: A Dynamic Panel Approach

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ABSTRACT

This paper investigates the effect of remittance inflows on real exchange rates in sub-Saharan Africa (SSA) using annual data from 1980 to 2008 for 34 countries, the method of moments estimator developed by Arellano and Bover (1995) and the feasible generalized least squares estimator developed by Parks (1967) and Kmenta (1986). We find that when cross-sectional dependence and individual effects are controlled for, remittances to sub-Saharan Africa as a whole increase the underlying real exchange rates of recipient countries. However, this real exchange rate appreciation is mitigated by monetary policy interventions and the direction of fiscal expenditures towards tradable goods. Thus, the real exchange rate appreciation does not lead to the loss of export competitiveness or a worsening of the trade deficit in the countries in the panel.

KEY WORDS:

Dutch disease; remittances; real exchange rate; sub-Saharan Africa

JEL Classification:

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Introduction

Remittances to developing countries have reached significant levels over the last two decades. As of the end of 2008, global remittances totaled thrice the value of official development assistance to developing countries and exceeded 10% of GDP in 23 recipient countries worldwide (Mohapatra et al., 2009). Remittances to sub-Saharan Africa steadily increased from 1.4 billion US dollars in 1980 to 21.3 billion US dollars in 2008, approximately 2.2% of regional

GDP (World Bank, 2008). It is further estimated that an additional amount equal to approximately 50% of formal inflows arrives in sub-Saharan Africa through informal channels (International Monetary Fund [IMF], 2006).

Research has demonstrated that significant increases in foreign inflows such as remittances could cause the underlying real exchange rate of the recipient economy to appreciate, adversely affecting export competitiveness and, consequently, the trade deficit (Corden and Neary, 1982). This would further result in the contraction of the tradable sector of the recipient economy, leading to a decline in the manufacturing and production of other tradable goods. Additionally, an increase in remittances - *ceteris paribus* - increases the disposable incomes of recipient households, leading to

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an increase in aggregate demand. This spending effect results in higher relative prices of non-tradable goods (i.e., goods produced and consumed domestically that are not close substitutes for imported or exported goods and services), as the prices of tradable goods (i.e., goods that are traded internationally and obey the law of one price or an appropriate relative pricing) are assumed to be exogenously given (Acosta, Lartey, & Mandelman, 2007). The higher prices of non-tradable goods lead to an expansion of the non-tradable sector. Assuming that resources are perfectly mobile, there could be a reallocation of resources from the tradable to the non-tradable sector. In addition to this reallocation of resources, remittance-receiving households are also known to occasionally reduce labor supply (Amuedo-Dorantes & Pozo, 2004). Assuming that resources are fully utilized, this could increase the marginal cost of labor in the tradable sector, leading to an increase in production costs and a further contraction of the tradable sector (Acosta et al., 2007). These adverse effects of an increase in foreign inflows (in this case, remittances) on the real exchange rate, the tradable sector, loss of export competitiveness, and consequently, the trade deficit are referred to as the Dutch disease effect of remittance inflows (Corden & Neary, 1982). However, this is based on the assumption that households primarily spend remittances on non-traded goods. The Dutch disease effect of foreign inflows could also be influenced by other fundamental determinants of real exchange that depreciate the real exchange rate, thereby mitigating the appreciating effect of foreign inflows on the real exchange rate (Edwards 1989; 1994; Montiel, 1999). Furthermore, in certain countries, a specific policy stance by monetary authorities and conditionalities on development assistance have also been found to mitigate the natural macroeconomic transmission mechanism of foreign inflows (Nwachukwu, 2008; Oomes, 2008).

The impact of foreign inflows on the real exchange rate has been found to vary from region to region. In a study on foreign aid and the real exchange rate in 12 francophone West African countries, Quattara and Strobl (2004) found that foreign aid flows do not generate Dutch disease effects. Similar results were obtained by Ogun (1995) for Nigeria and Nyoni (1998) for Tanzania. On the contrary, Elbadawi (1999), in a study of 62 developing countries, and White and

Wignaraja (1992), for Sri Lanka, reported that aid flows appreciated the real exchange rates of the recipient countries in their study. Conflicting results were also observed in studies of foreign aid and the real exchange rate in Ghana. While Sackey (2001) found no appreciating effect on the real exchange rate, Opoku-Afari, Morrissey and Lloyd (2004) found the opposite and support for the Dutch disease theory. Using annual data on six Central American countries from 1985 to 2004, Izquierdo and Montiel (2006) found the real exchange rate to be relatively stable despite increased remittance inflows. In other cases such as the Euro-Mediterranean region, remittance inflows appreciated the real exchange rate but did not result in a deterioration of the current account balance, although exports suffered to some extent (Oomes, 2008). These disparities in findings by those authors (Nyoni, 1998; Quattara & Strobl, 2004; Ogun, 1995; Sackey, 2001), on one hand, who found no effect of foreign inflows on the international competitiveness of recipient countries and those (Elbadawi, 1999; Izquierdo and Montiel, 2006; Opoku-Afari et al., 2004; White & Wignaraja, 1992), on the other hand, who found that foreign inflows adversely affected the international competitiveness of recipient countries merit the need to ascertain the precise impact of remittance inflows on the international competitiveness of recipient countries in sub-Saharan Africa.

The objective of this paper is to examine the impact of remittances on the real exchange rate using annual data from 1980 to 2008 for 34 sub-Saharan African (SSA) countries. Do remittances appreciate the real exchange rate of the recipient countries in the panel? If yes, does it lead to a loss of export competitiveness and a deterioration of the trade deficit? If no, is this because the appreciating effect is mitigated by fundamental determinants of the real exchange rate or an intervention by monetary authorities in pursuit of a specific monetary policy objective, such as maintaining export competitiveness or a sustainable current account deficit? We adopt the theoretical framework of Montiel (1999), which states that the real exchange rate is an endogenous variable and is in equilibrium when it is simultaneously consistent with internal and external balances and conditional on long-run fundamentals (sustainable values of exogenous and policy variables). Internal balance refers to the situation in which the non-tradable goods market clears in the

current period and is expected to be in equilibrium in the future (Montiel 1999). Thus, assuming initial internal balance equilibrium, an increase in private spending creates excess demand for non-tradable goods at the initial exchange rate. An appreciation of the real exchange rate would then be required to restore equilibrium. This would - *ceteris paribus* - lead to an increase in the supply of non-tradable goods and an increase in demand for tradable goods. The external balance, however, is defined as the current account balance that is consistent with sustainable long-run capital inflows (Montiel, 1999). This is determined by the domestic output of traded goods net of domestic consumption, plus net aid flows, less the cost of foreign debt. From the perspective of initial external balance equilibrium, an increase in private spending generates a current account deficit at the initial exchange rate. A real depreciation would therefore be required in this case to restore equilibrium. This leads to an increase in the supply of tradable goods and an increase in the demand for non-tradable goods.

The Montiel (1999) model states that factors affecting internal and external balance would also cause changes in the long-run equilibrium real exchange rate. These factors include fiscal policy, international transfers, terms of trade, total factor productivity, international financial conditions and commercial policy (see Montiel, 1999 and 2003, for complete details on the theoretical framework).

Most studies on the impact of foreign inflows on the real exchange rate in sub-Saharan Africa have primarily focused on official development assistance and devoted scant attention to remittances. Second, most of them have examined specific countries in sub-Saharan Africa such as Tanzania (Nyoni, 1998), Nigeria (Ogun, 1995), Ghana (Opoku-Afari et al., 2004; Sackey 2001) and rarely sub-regions within sub-Saharan Africa such as francophone West Africa (Ouattara & Strobl, 2004) or sub-Saharan Africa as a whole (Nwachukwu, 2008). This paper therefore contributes to the limited literature on remittance inflows into Sub-Saharan Africa and their effect on the real exchange rate, using annual data on 34 sub-Saharan African countries from 1980 to 2008 and dynamic panel estimation techniques, namely the feasible generalized least squares (FGLS) approach developed by Parks (1967) and Kmenta (1986) and the two-step system generalized methods of moments (GMM) proposed by Arellano and Bover

(1995). This paper differs from most previous work by testing for cross-sectional dependence in the error term among the countries in the panel using the Pesaran (2004) CD test and controlling for it using appropriate dynamic panel estimation techniques.

In the context of remittances, cross-sectional dependence is caused by the spillover effect of remittances across borders in the sub-Saharan African region. African countries have migrants from neighboring countries who receive remittances from their relatives in the diaspora. Labor mobility (both skilled and unskilled) and cross-border trade are highly prevalent within the region. This indicates that non-recipient households also benefit from remittance inflows through labor income and recipient households' demand for goods and services (Durrand, Parrado, & Massey, 1996). These factors (intra-African migration, labor mobility, cross-border trade and the multiplier effect of remittance flows) create spillover effects that have not been adequately addressed in the remittances literature on sub-Saharan Africa.

By addressing this gap in the sub-Saharan Africa remittances literature, we also control for a major assumption of panel data estimations, namely, cross-sectional independence. In the presence of cross-sectional dependence of the error term, methods employing instrumental variable estimations and the generalized method of moments (GMM) - such as those proposed by Anderson and Hsiao (1981), Arellano and Bond (1991), and Blundell and Bond (1998) - yield inconsistent estimates. Such inconsistency is exacerbated if one employs a short dynamic panel data model with a fixed T (the panel time dimension) as N (the cross-sectional dimension) grows large (Sarafidis & Robertson, 2009).

The remainder of this paper is structured as follows: section 2 describes the data and methodology, section 3 discusses empirical results, and section 4 contains the conclusions, policy recommendations and suggestions for future research.

Data and methodology

Table 1 below details the data series used and the measurement thereof. Following Montiel (1999, 2003), the fundamental determinants of the real exchange rate used in this paper include international transfers (remittances), fiscal expenditures (government final expenditure on goods and services), terms of trade, current account openness, international financial conditions

Table 1. Sources and definition of variables

	Variable	Source	Definition
rer	Real exchange rate	IMF, World Bank	The real exchange rate is measured as the product of the nominal bilateral exchange rate with the USD and the ratio of the US wholesale price index to domestic prices (GDP deflator, 2005 = 100) in each country.
rem	Remittances as a percentage of GDP (International transfers)	World Bank	Workers' remittances as a percentage of GDP in current prices (USD Millions).
fp	Fiscal policy	World Bank	Government final consumption as a percentage of GDP in SSA countries (a proxy for the direction of fiscal expenditure).
open	Current account openness	Penn World Table PWT 7.0	The ratio of the sum of exports and imports of goods and services to GDP in SSA countries.
idif	Interest rate differential	IMF, World Bank	Differential between the deposit interest rates in SSA countries and the USA.
m2	Monetary policy effects	IMF, World Bank	Quasi money as a percentage of GDP (a measure of monetary policy interventions).
tot	Terms of trade	World Bank	Ratio of export prices to import prices for SSA countries.

(interest rate differential between sub-Saharan African countries and the USA) and quasi money (M2) as a percentage of GDP, as a proxy for monetary policy effects.

The USA is employed as a “representative” host country for a number of reasons. The dominant destinations of migrants from the 34 sub-Saharan African (SSA) countries in this panel are UK, USA and Canada (for the Anglophone SSA countries); France and Belgium (for the Francophone SSA countries); and South Africa (for the SADC countries). An alternative approach would be to construct an index based on the weights of the GDP per capita of these six countries and use them to represent host country income/economic conditions. However, the annual GDP per capita measures of these six countries are highly positively correlated (above 0.98) due to the impact of global business cycle trends and global economic integration among these six countries. A graphical analysis of the GDP per capita of the six countries indicates that the GDP per capita of the other countries closely mirrors trends in the GDP per capita of the USA. The USA, Canada, France and Belgium are highly linked through EU-USA business cycle linkages, while the EU is the largest export destination for South Africa. Furthermore, the USA is a major trading partner for most of the countries in the panel. Thus, most African currencies are pegged to the US dollar (USD) in foreign exchange reserves. The USA also represents

the largest source of capital flows to African countries. Consequently, the USA has a strong impact on real exchange rate/interest rate differential dynamics in the countries in this panel. For these reasons, we use the largest/strongest economy of these dominant migration destinations (the USA) as a “**representative**” host country. It is also the only country where a “**representative migrant**” (not all or most of the migrants) from each of the 34 countries can be found.

Data on all variables for the sub-Saharan African countries in the panel, including Benin, Burundi, Botswana, Burkina Faso, Cameroun, Cape Verde, Comoros, Cote D'Ivoire, Ethiopia, Gabon, The Gambia, Ghana, Guinea, Guinea-Bissau, Kenya, Lesotho, Madagascar, Malawi, Mali, Mauritius, Mauritania, Mozambique, Niger, Nigeria, Rwanda, Senegal, the Seychelles, Sierra Leone, Swaziland, South Africa, Tanzania, Togo, Uganda, and Zambia are obtained from the World Development Indicators of the World Bank, complemented with data from the International Monetary Fund.

Table 2 details the relationship between the variables employed and the real exchange rate. This forms the basis of our a priori expectations.

2.1 Model specification and estimation technique

Cross-correlation analysis reveals that the real exchange rate exhibits strong persistence behavior. This warrants

Table 2. A priori expectations. Dependent variable: rer

Variable	Sign	Inference
rem	Positive/ negative	Remittances improve the foreign reserve position of recipient countries, which should appreciate the domestic currency (negative relationship). If remittances were spent on tradables, then it would have a depreciating effect on the real exchange rate (positive relationship).
fp	Positive/ negative	Fiscal expenditures on traded goods would have a depreciating effect on the real exchange rate (positive relationship). If they target non-traded goods, then they would have an (negative) appreciating effect on the real exchange rate.
tot	Positive/ negative	An export-dominated terms of trade would appreciate the real exchange rate (negative relationship), while an import-dominated terms of trade would depreciate the real exchange rate (positive relationship).
open	Positive/ negative	An export-dominated foreign sector would appreciate the real exchange rate (negative relationship); an import-dominated foreign sector would depreciate the real exchange rate (positive relationship).
idif	Negative	A positive interest rate differential should attract foreign inflows that should appreciate the real exchange rate (negative relationship).
m2	Negative/ positive	Monetary policy interventions (sterilization) designed to depreciate the real exchange rate would have a positive relationship with the real exchange rate and a negative relationship if it is intended to appreciate the real exchange rate.

employing a dynamic panel data model that includes one or more lags of the dependent variable. We specify a two-way error component model based on the heterogeneity among the 34 countries in the panel and various developments at different time periods that significantly affected exchange rate developments in sub-Saharan African during the sample period. These time effects include nominal devaluations by countries, oil price shocks and political instability at various time periods in different countries, all of which had varying effects on the nominal and, consequently, real exchange rates of the countries in the panel. These developments, especially political conflicts, also had spillover effects within the region. The model is therefore specified in (1) as

$$y_{it} = \delta y_{i,t-1} + X'_{it}\beta + \mu_i + \lambda_t + v_{it} \quad (1)$$

where y_{it} is an $NT \times 1$ vector of dependent and endogenous variables. X'_{it} represents an $NT \times k$ vector of independent regressors other than the lag of the dependent variable, β denotes a $k \times m$ vector of slope coefficients, μ_i represents country-specific effects, λ_t time effects and v_{it} the idiosyncratic error term. Equation (1) is based on the assumption of exogeneity of the regressors. However, by construction, the lag of the dependent variable $y_{i,t-1}$ is

correlated with the fixed-effects μ_i error term. Additionally, the terms of trade variable is correlated with the error term. Consequently, the Hausman specification test rejects the null of the exogeneity of all right-hand-side regressors in (1). A general LM-test for heteroskedasticity (Greene, 2003) indicates the existence of heteroskedasticity in the specification. The modified Wald test rejects the null of group-wise homoskedasticity, implying a non-constant variance across cross-sections. However, it is known to have very low power in the context of fixed effects when $N > T$ (Greene, 2003). This test is therefore not reported but controlled for in the analysis. Table 3 details the results of the initial diagnostic tests performed on the pooled ordinary least squares (OLS) and fixed-effects (FE) models.

In addition to the presence of endogeneity, the two-way fixed-effect model specified in (1) also suffers from first-order serial correlation. Furthermore, a test for cross-sectional dependence of the error terms using the Pesaran (2004) CD test rejects the null of cross-sectional independence at the 1% level of significance, with an average cross-sectional correlation coefficient of 0.42.

To establish the order of integration of the dataset, we employ the Pesaran (2007) cross-sectionally adjusted

Table 3. Initial diagnostic tests

Test	Test statistic	Critical value	Inference
Joint validity of cross-sectional individual effects $H_0: \mu_1 = \mu_2 \dots \mu_{N-1} = 0$ $H_A: \text{Not all equal to } 0$	F = 15.12	F(0.05, 34, 939) = 1.442	Cross-section-specific effects are valid.
Joint validity of time (period) effects $H_0: \lambda_1 = \dots \lambda_{T-1} = 0$ $H_A: \text{Not all equal to } 0$	F = 44.51	F(0.05, 27, 947) = 1.498	Time effects are valid. The error term takes a two-way error component form.
Heteroskedasticity $H_0: \sigma_i^2 = \sigma^2$ $H_A: \text{Not equal for all } i$	LM = 817.59	$\chi^2_{(34)}$ 48.60	The variance of the error term is not constant. Heteroskedasticity is present.
Hausman specification test $H_0: E(u_{it} X_{it}) = 0$ $H_A: E(u_{it} X_{it}) \neq 0$	$m_3 = 160.11$	$\chi^2_{(7)}$ 12.60	Regressors not exogenous.
Pesaran (2004) CD test for cross-sectional dependence $H_0: \text{corr}(u_{it}, u_{jt}) = 0 \text{ for } i \neq j$ $H_A: \text{corr}(u_{it}, u_{jt}) = 0 \text{ for some } i \neq j$	CD = 5.2 (0.42)	Prob = 0.000	Cross-sections are interdependent.

Dickey-Fuller (CADF) Test. The Pesaran (2007) CADF test is well suited to heterogeneous panel data with cross-sectional dependence in the errors, as is the case in this study. It assumes a null hypothesis of non-stationarity of the data series. In the Pesaran (2007) CADF test, the cross-sectional dependence is eliminated by augmenting Dickey-Fuller (or Augmented Dickey-Fuller) regressions with the cross-sectional averages of lagged levels and first-differences of the individual series (CADF statistics). The results of unit root tests are reported in Table 4. The variables *rer*, *m2*, *rem* and *open* are found to be non-stationary I(1), while *fp*, *tot* and *idif* are stationary.

The results of the initial diagnostic tests warrant the use of an estimation technique that preserves homoskedasticity, controls for cross-sectional dependence and preserves the orthogonality between transformed variables and lagged regressors (Arellano and Bover, 1995). Two estimation techniques meet these criteria, namely the feasible generalized least squares (FGLS) technique developed by Parks (1967) and Kmenta (1986) and the two-step system generalized method of moments (GMM) proposed by Arellano and Bover (1995).

The Parks and Kmenta FGLS estimation technique is perfectly suited to data with individual effects, group-wise heteroskedasticity, serial correlation and cross-sectional dependence in the error term (Hicks, 1994; Kmenta, 1986). The FGLS estimation technique is suitable whether the individual effects are fixed over time and cross-sections or are normally distributed random variables. It involves two sequential transformations. First, it eliminates serial correlation in the errors and then the contemporaneous correlation of the errors across cross-sections, which also automatically corrects for panel heteroskedasticity (Beck & Katz, 1995). This is achieved by initially estimating a pooled model. Residuals from this initial estimation are used to estimate the unit-specific serial correlation of the errors, which are then used to transform the model into one with serially independent errors. The residuals from this estimation are then used to estimate the contemporaneous correlation of the errors, and the data is again transformed to allow for an estimation without any complications in the errors. This process yields consistent estimators of the elements of the variance-covariance matrix, which then yields the desired

Table 4. Pesaran (2007) Cross-sectional Augmented Dickey-Fuller (CADF) unit root test

	rer	fp	tot	idif	m2	rem	open
Pesaran (2007)							
t-stat							
Levels	1.320	-2.113**	-1.974*	-2.017**	-1.574	-1.715	-1.689
[P-value]	[0.99]	[0.017]	[0.09]	[0.06]	[0.869]	[0.67]	[0.67]
Differences	-3.942***	--	--	--	-3.258***	-3.499***	-3.644***
[P-value]	[0.00]				[0.00]	[0.00]	[0.00]

Note: ***/**/* denote 1/5/10 percent levels of significance, respectively. Statistical insignificance implies failing to reject the null of non-stationarity; p-values in [. Critical values at 1/5/10 percent available in Pesaran (2007) and in test results output in STATA 11.

coefficient estimates and their respective standard errors (Kmenta, 1986). The FGLS estimation is, however, known to lose some degree of efficiency when there is endogeneity from regressors other than the lagged dependent variable (Kmenta, 1986). Thus to ensure robustness, we also employ the two-step system GMM estimation technique of Arellano and Bover (1995).

In the two-step system GMM, the endogeneity problem is addressed by time demeaning the data to remove time effects. Time demeaning the data is equivalent to including common time effects to capture “common trends” in the variation of the dependent variable across cross-sections. Time demeaning the data reduces cross-sectional dependence “but only to a certain extent” (De Hoyos & Sarafidis, 2006). As we have a moderate level of cross-sectional dependence in this paper (cross-correlation coefficient of 0.42), time demeaning the data prior to the GMM estimation should partially or sufficiently control for it. The cross-section specific effects are then eliminated using forward orthogonal deviations, thereby making it possible to use one-period lags of the regressors as valid instruments, as they are not correlated with the transformed error term (Love & Zichinno, 2006). We estimate the two-step system GMM while collapsing the lag range to avoid instrument proliferation and apply the Windmeijer (2005) correction for heteroskedasticity to ensure robust standard errors. Another advantage of this approach is that it is more resilient to missing data. It is computable for all observations except for the last for each cross-section, thereby minimizing data loss (Roodman, 2006).

Empirical results

The FGLS and the two-step system GMM estimations yield similar results. Table 5 details the empirical results. The FGLS estimation specified that the errors of the panels are correlated. The two-step system GMM estimation involves forward orthogonal deviations instead of differencing (Arellano and Bover, 1995). The OLS results are significantly different from the FGLS and two-step system GMM results due to violations of the assumptions of the classical linear regression model, as reported in the results of initial diagnostics in Table 3.

As expected, the real exchange rate exhibits strong persistence behavior that is significant at the 1% level. The coefficient of remittance inflows is negatively signed and statistically significant at the 1% level. This means that remittances on average have an appreciating effect on the real exchange rates of the recipient sub-Saharan African countries in the panel.

Fiscal expenditure is positively signed and statistically significant at the 1% level. This indicates that government expenditure is more targeted towards traded goods, with the economy requiring an exchange rate depreciation to restore external balance. The coefficient of terms of trade is negatively signed and statistically significant at the 1% level, indicating an appreciating effect on the real exchange rate. This denotes that the income effect dominates the substitution effect of an improvement in the terms of trade, requiring an appreciation of the real exchange rate to restore external balance. Current account openness is also negatively signed and statistically significant at the 1% level. This indicates an export-dominated foreign

Table 5. Empirical results: FGLS and two-step system GMM. Dependent variable: *rer*

Variable	OLS	FGLS	Two-step system GMM
<i>rer(-1)</i>	0.90*** (32.14)	0.79*** (42.80)	0.75*** (32.38)
<i>rem</i>	0.28** (2.12)	-3.05*** (-3.03)	-3.20*** (-2.22)
<i>fp</i>	-1.84*** (-3.58)	1.04*** (8.49)	1.07*** (6.68)
<i>tot</i>	-0.08** (-2.00)	-0.72*** (-5.99)	-0.99*** (-6.17)
<i>open</i>	-0.11** (-2.14)	-0.64*** (-5.53)	-0.93*** (-3.23)
<i>idif</i>	0.08 (1.33)	0.81*** (4.32)	0.71*** (9.83)
Δm^2	-0.52*** (-3.46)	1.21*** (2.68)	2.88*** (3.32)
Adjusted R ²	0.98		
A Bond test for second-order serial correlation			Prob > z = 0.29
Hansen test for over-identification			Prob $\chi^2 > = .36$
Diff. in Hansen test for exogeneity of instrument set			Prob $\chi^2 > = .37$

Notes: */**/** denote 10/5/1 percent levels of significance, respectively. T-statistics are in parentheses. When performing the Hansen test for over-identification, the "collapse" option in STATA was used to reduce the lag range and avoid instrument proliferation, in conjunction with the Windmeijer (2005) correction for robust standard errors.

sector on average and hence an appreciating effect on the real exchange rate. Contrary to *a priori* expectations, the interest rate differential is positively signed and statistically significant at the 1% level, which denotes a depreciating effect on the real exchange rate of a recipient country following an increase in the home country interest rate relative to the host country interest rate. This result may be attributed to the fact that remittances are less responsive to changes in the interest rate differential, over and above the fact that it only constitutes a portion of total inflows to any recipient country. The monetary policy variable is positively signed and statistically significant at the 1% level. This denotes that monetary policy is formulated to keep

the exchange rate low probably to ensure a sustainable current account deficit. This provides an indication of the mitigating effect of monetary policy interventions such as sterilization on the appreciation of the real exchange rate. Such monetary policy intervention by countries is generally intended to maintain a real exchange rate that ensures export competitiveness and a sustainable current account deficit.

The two-step system GMM estimation satisfies all post-estimation diagnostic requirements. The Arellano and Bond (1991) test for second-order serial correlation fails to reject the null of no autocorrelation. The Hansen (1982) test for over-identification fails to reject the null that the over-identification restrictions

are valid, while the Difference-in-Hansen test also fails to reject the null that the instrument subset is strictly exogenous.

Conclusions, policy recommendations and future research

The empirical results demonstrate that when cross-sectional dependence and individual effects are controlled for, remittance inflows on average appreciate the underlying exchange rate of the recipient economies in this study. This is consistent with the Dutch-disease theory of Corden and Neary (1982). However, this appreciating effect of remittance inflows on the real exchange rate is mitigated by monetary policy interventions and the direction of fiscal expenditure.

Monetary policy is formulated to maintain a low real exchange rate, despite the steady increases in the rate of inflation observed for the countries in the panel. This prevents remittance inflows from exerting their natural transmission effect on the real exchange rate. It then implies that the nominal exchange rate is managed to some extent in most of the countries in the panel. This aligns with the findings of Oomes (2008) on Armenia and Nwachukwu (2008) on sub-Saharan Africa, who cite policy interventions as the mitigating factor for the appreciating effect of foreign inflows on the real exchange rate.

In addition to the monetary policy stance, the direction of fiscal expenditure increases the likelihood that fiscal expenditure is more targeted towards tradables than non-tradables, hence its depreciating effect on the real exchange rate. This is understandable due to the rather low levels of production capacity in the sub-Saharan African countries in the panel and hence these economies' substantial dependence on imports. Consequently, a deteriorating current account deficit is more likely due to over-dependence on imports than a loss of export competitiveness resulting from the impact of remittance inflows on the real exchange rate.

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